

AN L^1 ERGODIC THEOREM FOR SPARSE RANDOM SUBSEQUENCES

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ABSTRACT. We prove an L^1 subsequence ergodic theorem for sequences chosen by independent random selector variables, thereby showing the existence of universally L^1 -good sequences nearly as sparse as the set of squares. In the process, we prove that a certain deterministic condition implies a weak maximal inequality for a sequence of ℓ^1 convolution operators (Prop. 3.1).

1. Introduction

Let (X, \mathcal{F}, m) be a non-atomic probability space and T a measure-preserving transformation on X ; we call (X, \mathcal{F}, m, T) a dynamical system. For a sequence of integers $\mathbf{n} = \{n_k\}$ and any $f \in L^1(X)$, we may define the subsequence average

$$A_N^{(\mathbf{n})} f(x) := \frac{1}{N} \sum_{k=1}^N f(T^{n_k} x).$$

Given a sequence \mathbf{n} , a major question is for which $1 \leq p \leq \infty$ and which (X, \mathcal{F}, m, T) we have convergence of various sorts for all $f \in L^p(X)$. An important definition along these lines is as follows:

Definition A sequence of integers $\mathbf{n} = \{n_k\}$ is *universally L^p -good* if for every dynamical system (X, \mathcal{F}, m, T) and every $f \in L^p(X, m)$, $\lim_{N \rightarrow \infty} A_N^{(\mathbf{n})} f(x)$ exists for almost every $x \in X$.

Birkhoff's Ergodic Theorem asserts, for instance, that the sequence $n_k = k$ is universally L^1 -good. On the other extreme, the sequence $n_k = 2^k$ is not even universally L^∞ -good (lacunary sequences are bad for convergence of ergodic averages in various strong ways, see for example [10] or [1]). Between these extrema lie many results on the existence of universally L^p -good sequences of various sorts, beginning with Bourgain's celebrated result [5] that $n_k = k^2$ is universally L^2 -good; see [6] and [3] for extensions of this result to other sequences.

The most restrictive case $p = 1$ is more subtle than the others. A surprising illustration of the difference is the recent result of Buczolich and Mauldin that $n_k = k^2$ is not universally L^1 -good [8]. Positive results in L^1 have been difficult to come by, particularly for sequences which are sparse in \mathbb{N} .

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Universally L^1 -good sequences of density 0 had long been known to exist, but these were sparse block sequences, which consist of large 'blocks' of consecutive integers, separated by wide gaps. Bellow and Losert [2] showed that for any $F : \mathbb{N} \rightarrow \mathbb{R}^+$, there exists a universally L^1 -good block sequence $\{n_k\}$ with $n_k \geq F(k)$. To distinguish such block sequences from more uniformly distributed ones, we recall the notion of Banach density:

Definition A sequence of positive integers $\{n_k\}$ has Banach density c if

$$\limsup_{m \rightarrow \infty} \sup_N \frac{|\{n_k \in [N, N+m]\}|}{m} = c.$$

Note that block sequences with arbitrarily large block lengths have Banach density 1 (the sequences in [2] are all of this sort). The first example of a Banach density 0 universally L^1 -good sequence was constructed by Buczolic [7], and Urban and Zienkiewicz [13] subsequently proved that the sequence $[k^a]$ for $1 < a < 1 + \frac{1}{1000}$ is universally L^1 -good.

Bourgain [5] showed that certain sparse random sequences were universally L^p -good with probability 1 for all $p > 1$. These sequences are generated as follows: given a decreasing sequence of probabilities $\{\tau_j : j \in \mathbb{N}\}$, let $\{\xi_j : j \in \mathbb{N}\}$ be independent random variables on a probability space Ω with $\mathbb{P}(\xi_j = 1) = \tau_j$, $\mathbb{P}(\xi_j = 0) = 1 - \tau_j$. Then for each $\omega \in \Omega$, define a random sequence by taking the set $\{n : \xi_n(\omega) = 1\}$ in increasing order. (For $\alpha > 0$ and $\tau_j = O(j^{-\alpha})$, these sequences have Banach density 0 with probability 1; see Prop. 4.3 of this paper.)

In this paper, we apply a construction of [13] to these random sequences and achieve the following L^1 result:

Theorem 1.1. *Let $0 < \alpha < 1/2$, and let ξ_n be independent selector variables on Ω with $\mathbb{P}(\xi_n = 1) = n^{-\alpha}$. Then there exists a set $\Omega' \subset \Omega$ of probability 1 such that for every $\omega \in \Omega'$, $\{n : \xi_n(\omega) = 1\}$ is universally L^1 -good.*

Thus we prove the existence of universally L^1 -good sequences which grow more rapidly than the ones obtained in [13] or [7], and which grow uniformly as compared to the sparse block sequences of [2]. In particular, with probability 1 these sequences have $n_k = \Theta(k^{1/(1-\alpha)})$ (that is, $c_\omega k^{1/(1-\alpha)} \leq n_k \leq C_\omega k^{1/(1-\alpha)}$), so Theorem 1.1 applies to random sequences nearly as sparse as the sequence of squares.

Our method is as follows: in Section 2 we define our notation and reduce the problem (by transference) to one of proving a weak maximal inequality on \mathbb{Z} for convolutions with a series of random $\ell^1(\mathbb{Z})$ functions $\mu_n^{(\omega)}$. In Section 3, we use the framework of [13] to prove this inequality under an assumption about the convolutions of $\mu_n^{(\omega)}$ with their reflections about the origin; and in Section 4, we establish that with probability 1, these random functions do indeed satisfy that assumption.

2. Definitions, and Reduction to a Weak Maximal Inequality

Let $\{\tau_n : n \in \mathbb{N}\}$ be a nonincreasing sequence of probabilities. Let Ω be a probability space, and define independent mean τ_n Bernoulli random variables $\{\xi_n(\omega) : n \in \mathbb{N}\}$ on Ω ; that is, $\mathbb{P}(\xi_n = 1) = \tau_n$ and $\mathbb{P}(\xi_n = 0) = 1 - \tau_n$. Let

$$\beta(N) := \sum_{n=1}^N \tau_n.$$

Definition For a dynamical system (X, \mathcal{F}, m, T) and $f \in L^1(X)$, define the random average

$$A_N^{(\omega)} f(x) := \beta(N)^{-1} \sum_{n=1}^N \xi_n(\omega) f(T^n x)$$

and its $L^1(X)$ -valued expectation

$$\mathbb{E}_\omega A_N^{(\omega)} f(x) := \beta(N)^{-1} \sum_{n=1}^N \tau_n f(T^n x).$$

Remark $A_N^{(\omega)} f$ differs from the subsequence averages discussed before by the factor $\beta(N)^{-1} \sum_{n=1}^N \xi_n(\omega)$. However, if $\beta(N) \rightarrow \infty$, then $\beta(N)^{-1} \sum_{n=1}^N \xi_n(\omega) \rightarrow 1$ almost surely in Ω . This follows directly from the first Borel-Cantelli Lemma and from Chernoff's Inequality, which we will use elsewhere in this paper:

Theorem 2.1. *Let $\{X_n\}_{n=1}^N$ be independent random variables with $|X_n| \leq 1$ and $\mathbb{E}X_n = 0$. Let $X = \sum_{n=1}^N X_n$, and $\sigma^2 = \mathbf{Var} X = \mathbb{E}X^2$. Then for any $\lambda > 0$,*

$$\mathbb{P}(|X| \geq \lambda\sigma) \leq 2 \max(e^{-\lambda^2/4}, e^{-\lambda\sigma/2}).$$

Proof. This is Theorem 1.8 in [12], for example. □

We restrict to the set of full measure $\Omega_1 \subset \Omega$ on which $\beta(N)^{-1} \sum_{n=1}^N \xi_n(\omega) \rightarrow 1$.

The a.e. convergence of $A_N^{(\omega_0)} f(x)$ for every dynamical system (X, \mathcal{F}, m, T) and every $f \in L^p(X)$ is then equivalent to the statement that $\{j \in \mathbb{N} : \xi_j(\omega_0) = 1\}$ is universally L^p -good. We further remark that for a power law $\tau_n = n^{-\alpha}$, we have $N^{\alpha-1}\beta(N) \rightarrow C \in (0, \infty)$ for $\alpha < 1$.

By Bourgain's result in [5], there is a set $\Omega_2 \subset \Omega_1$ with $\mathbb{P}(\Omega_2) = 1$ such that for $\omega \in \Omega_2$ we have a.e. convergence of $A_N^{(\omega)} f$ for all $f \in L^2(X)$, which is dense in $L^1(X)$. Theorem 1.1 thus reduces to proving on a set of probability 1 the weak maximal inequality

$$(2.1) \quad \left\| \sup_N |A_N^{(\omega)} f| \right\|_{1, \infty} \leq C_\omega \|f\|_1 \quad \forall f \in L^1(X).$$

As usual, it is enough to take this supremum over the dyadic subsequence $\{2^j : j \in \mathbb{N}\}$, since $\frac{\beta(2^{j+1})}{\beta(2^j)} \leq 2$ and thus $0 \leq A_N^{(\omega)} f \leq 2A_{2^{j+1}}^{(\omega)} f$ for $f \geq 0$ and $2^j \leq N < 2^{j+1}$. As in [4] and other papers, we can transfer this problem to the group algebra $\ell^1(\mathbb{Z})$. Namely, if we define the random $\ell^1(\mathbb{Z})$ functions

$$\begin{aligned} \mu_j^{(\omega)}(n) &:= \begin{cases} \beta(2^j)^{-1} \xi_n(\omega), & 1 \leq n \leq 2^j \\ 0 & \text{otherwise} \end{cases} \\ \mathbb{E}\mu_j(n) &:= \begin{cases} \beta(2^j)^{-1} \tau_n, & 1 \leq n \leq 2^j \\ 0 & \text{otherwise} \end{cases} \\ \nu_j^{(\omega)}(n) &:= \mu_j^{(\omega)}(n) - \mathbb{E}\mu_j^{(\omega)}(n), \end{aligned}$$

then $\mu_j^{(\omega)}$ and $\mathbb{E}\mu_j$ correspond to the operators $A_{2^j}^{(\omega)}$ and $\mathbb{E}_\omega A_{2^j}^{(\omega)}$, respectively. By a standard transference argument, it suffices to prove that with probability 1 in Ω ,

$$(2.2) \quad \left\| \sup_j |\varphi * \mu_j^{(\omega)}| \right\|_{1,\infty} \leq C_\omega \|\varphi\|_1 \quad \forall \varphi \in \ell^1(\mathbb{Z}).$$

We will use $\tilde{\mu}$ to denote the reflection of a function μ about the origin; as the adjoint of the operator given by convolution with μ is a convolution with $\tilde{\mu}$, this will be an important object. (It would be standard to use the notation μ^* , but this becomes unwieldy when using other superscripts as above.)

3. Calderon-Zygmund Argument

The proof of (2.2) uses a generalization of a deterministic argument from the paper by Urban and Zienkiewicz [13], related to a construction of Christ in [9]:

Proposition 3.1. *Let μ_j and ν_j be sequences of functions in $\ell^1(\mathbb{Z})$. Let $r_j := |\text{supp } \mu_j|$ and suppose $\text{supp } \nu_j \subset [-R_j, R_j]$. Assume there exist $\epsilon > 0$ and $A, A_0, A_1 < \infty$ such that $\sum_{j \leq k} r_j \leq A r_k \forall k \in \mathbb{N}$ and*

$$(3.1) \quad |\nu_j * \tilde{\nu}_j(x)| \leq A_0 r_j^{-1} \delta_0(x) + A_1 R_j^{-(1+\epsilon)}, \quad \forall x \in \mathbb{Z}.$$

If there exist $p \in (1, \infty]$ and $C, C_p < \infty$ such that for all $\varphi \in \ell^1(\mathbb{Z})$,

$$\left\| \sup_j \varphi * |\mu_j - \nu_j| \right\|_{1,\infty} \leq C \|\varphi\|_1 \quad \text{and} \quad \left\| \sup_j |\varphi * \mu_j| \right\|_{p,\infty} \leq C_p \|\varphi\|_p,$$

then there exists $C' < \infty$ such that

$$(3.2) \quad \left\| \sup_j |\varphi * \mu_j| \right\|_{1,\infty} \leq C' \|\varphi\|_1 \quad \forall \varphi \in \ell^1(\mathbb{Z}).$$

Proof. We will follow the argument in Section 3 of [13], which makes use of a Calderon-Zygmund type decomposition of φ depending on the index j . We begin with the standard decomposition at height $\lambda > 0$: $\varphi = g + b$, where

- $\|g\|_\infty \leq \lambda$
- $b = \sum_{(s,k) \in \mathcal{B}} b_{s,k}$ for some index set $\mathcal{B} \subset \mathbb{N} \times \mathbb{Z}$
- $b_{s,k}$ is supported on the dyadic cube $Q_{s,k} = [k2^s, (k+1)2^s) \cap \mathbb{Z}$
- $\{Q_{s,k} : (s,k) \in \mathcal{B}\}$ is a disjoint collection
- $\|b_{s,k}\|_1 \leq \lambda |Q_{s,k}| = \lambda 2^s$

- $\sum_{(s,k) \in \mathcal{B}} |Q_{s,k}| \leq \frac{C}{\lambda} \|\varphi\|_1$ (C independent of φ and λ).

Let $b_s = \sum_k b_{s,k}$. We will divide $\sum_s b_s$ into two parts, splitting at the index

$$s(j) := \min\{s : 2^s \geq R_j\}.$$

We begin by noting

$$\begin{aligned} \{x : \sup_j |\varphi * \mu_j(x)| > 4\lambda\} &\subset \{ \sup_j |g * \mu_j| > \lambda \} \cup \{ \sup_j |b * (\mu_j - \nu_j)| > \lambda \} \\ &\cup \{ \sup_j | \sum_{s=s(j)}^\infty b_s * \nu_j | > \lambda \} \cup \{ \sup_j | \sum_{s=0}^{s(j)-1} b_s * \nu_j | > \lambda \} \\ &= E_1 \cup E_2 \cup E_3 \cup E_4. \end{aligned}$$

By the weak (p, p) inequality (if $p < \infty$), $|E_1| \leq C\lambda^{-p} \|g\|_p^p \leq C\lambda^{-p} \|g\|_\infty^{p-1} \|g\|_1 \leq C\lambda^{-1} \|\varphi\|_1$; if $p = \infty$, consider instead that $\{x : \sup_j |g * \mu_j(x)| > C_\infty \lambda\} = \emptyset$ since $\|\sup_j |g * \mu_j|\|_\infty \leq C_\infty \|g\|_\infty \leq C_\infty \lambda$.

Next, $|b * (\mu_j - \nu_j)(x)| \leq |b * |\mu_j - \nu_j|(x)|$, so by the assumed weak $(1, 1)$ inequality,

$$|E_2| \leq |\{ \sup_j |b * |\mu_j - \nu_j| > \lambda \}| \leq \frac{C}{\lambda} \|b\|_1 \leq \frac{C}{\lambda} \|\varphi\|_1.$$

To bound $|E_3|$, note that for $s \geq s(j)$, $\text{supp}(b_{s,k} * \nu_j) \subset Q_{s,k} + [-R_j, R_j] \subset Q_{s,k}^*$, an expansion of $Q_{s,k}$ by a factor of 3. Thus

$$E_3 \subset \bigcup_j \bigcup_{k \in \mathbb{Z}, s \geq s(j)} \text{supp}(b_{s,k} * \nu_j) \subset \bigcup_{k \in \mathbb{Z}, s \geq s(j)} Q_{s,k}^*$$

and

$$|E_3| \leq \sum_{(s,k) \in \mathcal{B}} 3|Q_{s,k}| \leq \frac{C}{\lambda} \|\varphi\|_1.$$

We have thus reduced the problem to obtaining a bound on $|E_4|$. We will attempt this directly for heuristic purposes, and then modify our setup for the actual argument. By Chebyshev's Inequality,

$$\begin{aligned} |\{x : \sup_j | \sum_{s=0}^{s(j)-1} b_s * \nu_j(x) | > \lambda\}| &\leq \lambda^{-2} \sum_x \sup_j | \sum_{s=0}^{s(j)-1} b_s * \nu_j(x) |^2 \\ &\leq \lambda^{-2} \sum_j \left\| \sum_{s=0}^{s(j)-1} b_s * \nu_j \right\|_{\ell^2}^2 \\ &= \lambda^{-2} \sum_j \sum_{\substack{s_1, s_2 : \\ 0 \leq s_1, s_2 < s(j)}} \langle b_{s_1} * \nu_j, b_{s_2} * \nu_j \rangle_{\ell^2} \end{aligned}$$

and we will use our estimate on the convolution product $\nu_j * \tilde{\nu}_j$:

Lemma 3.2. *Let $f, g \in \ell^1$ such that $\sum_{x \in Q_{s(j),k}} |g(x)| \leq \lambda 2^{s(j)}$ for all k , and assume the ν_j satisfy (3.1). Then*

$$|\langle f * \nu_j, g * \nu_j \rangle| \leq A_0 r_j^{-1} |\langle f, g \rangle| + 10A_1 \lambda R_j^{-\epsilon} \|f\|_1.$$

Proof.

$$\begin{aligned} |\langle f * \nu_j, g * \nu_j \rangle| &= |\langle f * \nu_j * \tilde{\nu}_j, g \rangle| \\ &\leq A_0 r_j^{-1} |\langle f, g \rangle| + A_1 R_j^{-(1+\epsilon)} \|f\|_1 \|g\|_1. \end{aligned}$$

We let $f_k = f\chi(Q_{s(j),k})$ and $g_l = g\chi(Q_{s(j),l})$; note that $\|g_l\|_1 \leq \lambda 2^{s(j)} \leq 2\lambda R_j$. If $|k - l| > 2$, then $\langle f_k * \nu_j, g_l * \nu_j \rangle = 0$ as the supports are disjoint; thus

$$\begin{aligned} |\langle f * \nu_j, g * \nu_j \rangle| &\leq \sum_k \sum_{i=-2}^2 |\langle f_k * \nu_j, g_{k+i} * \nu_j \rangle| \\ &\leq \sum_k \sum_{i=-2}^2 A_0 r_j^{-1} |\langle f_k, g_{k+i} \rangle| + 2A_1 \lambda R_j^{-\epsilon} \|f_k\|_1 \\ &\leq A_0 r_j^{-1} |\langle f, g \rangle| + 10A_1 \lambda R_j^{-\epsilon} \|f\|_1. \end{aligned}$$

□

Therefore

$$\begin{aligned} |E_4| &\leq \lambda^{-2} \sum_j \sum_{\substack{s_1, s_2 : \\ 0 \leq s_1, s_2 < s(j)}} A_0 r_j^{-1} |\langle b_{s_1}, b_{s_2} \rangle| + 10A_1 \lambda R_j^{-\epsilon} \|b_{s_1}\|_1 \\ &\leq \lambda^{-2} \sum_j \sum_{0 \leq s < s(j)} A_0 r_j^{-1} \|b_s\|_2^2 + 10A_1 \lambda s(j) R_j^{-\epsilon} \|b_s\|_1 \\ &\leq A_0 \lambda^{-2} \sum_j r_j^{-1} \|b\|_2^2 + 10A_1 \lambda^{-1} \sum_j \log_2(2R_j) R_j^{-\epsilon} \|b\|_1. \end{aligned}$$

The assumption $\sum_{j \leq k} r_j \leq A r_k \forall k \in \mathbb{N}$ implies that r_j and R_j grow faster than any polynomial; thus the second term is $\leq \frac{C}{\lambda} \|\varphi\|_1$ as desired. The first term does not, however, give us that bound. We will therefore decompose these functions further.

For each j , we decompose $b_{s,k} = b_{s,k}^{(j)} + B_{s,k}^{(j)}$, where $b_{s,k}^{(j)} = b_{s,k} \chi(|b_{s,k}| > \lambda r_j)$. Define

$b_s^{(j)}, B_s^{(j)}, b_j^{(j)}, B_j^{(j)}$ by summing over one or both indices, respectively. Then

$$\begin{aligned} \left\{ \sup_j \left| \sum_{s=0}^{s(j)-1} b_s * \nu_j \right| > 3\lambda \right\} &\subset \left\{ \sup_j \left| \sum_{s=0}^{s(j)-1} b_s^{(j)} * (\nu_j - \mu_j) \right| > \lambda \right\} \\ &\cup \left\{ \sup_j \left| \sum_{s=0}^{s(j)-1} b_s^{(j)} * \mu_j \right| > \lambda \right\} \\ &\cup \left\{ \sup_j \left| \sum_{s=0}^{s(j)-1} B_s^{(j)} * \nu_j \right| > \lambda \right\} \\ &= E_5 \cup E_6 \cup E_7. \end{aligned}$$

We control E_5 just as we controlled E_2 , since $|b^{(j)}| \leq |b|$; and

$$\begin{aligned} |E_6| &\leq \sum_j |\{x : |b^{(j)} * \mu_j(x)| > 0\}| \leq \sum_j |\text{supp } \mu_j| \cdot |\{x : |b(x)| > \lambda r_j\}| \\ &= \sum_j r_j \sum_{k \geq j} |\{x : \lambda r_k < |b(x)| \leq \lambda r_{k+1}\}| \\ &= \sum_k |\{x : \lambda r_k < |b(x)| \leq \lambda r_{k+1}\}| \sum_{j \leq k} r_j \\ &\leq \frac{A}{\lambda} \sum_k \lambda r_k |\{x : \lambda r_k < |b(x)| \leq \lambda r_{k+1}\}|; \end{aligned}$$

now since this sum is a lower sum for $|b|$, we have $|E_6| \leq \frac{A}{\lambda} \|b\|_1 \leq \frac{C}{\lambda} \|\varphi\|_1$.

We proceed with E_7 just as we tried before, since Lemma 3.2 applies to the $B_s^{(j)}$ as well as to the b_s . We thus find

$$\begin{aligned} |E_7| &\leq A_0 \lambda^{-2} \sum_j r_j^{-1} \|B^{(j)}\|_2^2 + 10A_1 \lambda^{-1} \sum_j \log_2(2R_j) R_j^{-\epsilon} \|B^{(j)}\|_1 \\ &\leq A_0 \lambda^{-2} \sum_x \sum_j r_j^{-1} |B^{(j)}(x)|^2 + \frac{C}{\lambda} \|\varphi\|_1. \end{aligned}$$

Now $\sum_{j \leq k} r_j \leq A r_k \forall k \in \mathbb{N}$ implies $\exists N$ s.t. $r_{j+n} \geq 2r_j \forall j \in \mathbb{N}, n \geq N$, which implies $\sum_{j=k}^\infty r_j^{-1} \leq A' r_k^{-1}$; thus for each $x \in \mathbb{Z}$,

$$\sum_j r_j^{-1} |B^{(j)}(x)|^2 \leq \sum_{j: \lambda r_j \geq |b(x)|} r_j^{-1} |b(x)|^2 \leq A' \lambda |b(x)|$$

so $|E_7| \leq \frac{C}{\lambda} \|\varphi\|_1$ and the proof of Proposition 3.1 is complete. \square

4. Probabilistic Lemma, Conclusion of the Proof

Having established Proposition 3.1, it remains to show that the random measures $\mu_j^{(\omega)}$ and $\nu_j^{(\omega)}$ satisfy the assumptions with probability 1. Note first that

$r_j = |\text{supp } \mu_j^{(\omega)}| = \sum_{1 \leq n \leq 2^j} \xi_n(\omega) = \Theta(\beta(2^j)) = \Theta(2^{(1-\alpha)j})$ on Ω_1 , and $R_j = 2^{j+1}$. We will prove the bound (3.1) on $\nu_j^{(\omega)} * \tilde{\nu}_j^{(\omega)}$ via the following lemma:

Lemma 4.1. *Let $E \subset \mathbb{Z}$, and let $\{X_n\}_{n \in E}$ be independent random variables with $|X_n| \leq 1$ and $\mathbb{E}X_n = 0$. Assume that $\sum_{n \in E} (\mathbf{Var} X_n)^2 \geq 1$. Let X be the random ℓ^1 function $\sum_{n \in E} X_n \delta_n$, and let \mathbb{Z}^\times denote $\mathbb{Z} \setminus \{0\}$. Then for any $\theta > 0$,*

$$(4.1) \quad \mathbb{P} \left(\|X * \tilde{X}\|_{\ell^\infty(\mathbb{Z}^\times)} \geq \theta \left(\sum_{n=1}^N (\mathbf{Var} X_n)^2 \right)^{1/2} \right) \leq 4|E|^2 \max(e^{-\theta^2/16}, e^{-\theta/4}).$$

Proof. For $k \neq 0$,

$$X * \tilde{X}(k) = \sum_{n \in E \cap E-k} X_n X_{n+k} = \sum_{n \in E} Y_n$$

where $\mathbb{E}Y_n = 0$ and $|Y_n| \leq 1$ (of course $Y_n \equiv 0$ if $n + k \notin E$). We want to apply Chernoff's Inequality (Theorem 2.1), but the Y_n are not independent.

However, we can easily partition E into two subsets E_1 and E_2 such that $E_i \cap (E_i - k) = \emptyset$ for each i ; then within each E_i , the Y_n depend on distinct independent random variables, so they are independent.

Now $\sum_{n \in E_i} Y_n$ has variance $\sigma_i^2 = \sum_{n \in E_i} \mathbf{Var} X_n \mathbf{Var} X_{n+k} \leq \sum_{n \in E} (\mathbf{Var} X_n)^2$ by Hölder's Inequality. Chernoff's Inequality states that for any $\lambda > 0$,

$$\mathbb{P} \left(\left| \sum_{n \in E_i} Y_n \right| \geq \lambda \sigma_i \right) \leq 2 \max(e^{-\lambda^2/4}, e^{-\lambda \sigma_i/2}).$$

Take $\lambda_i = \theta \sigma_i^{-1} (\sum_{n \in E} (\mathbf{Var} X_n)^2)^{1/2}$; then $\lambda_i \sigma_i = \theta (\sum_{n \in E} (\mathbf{Var} X_n)^2)^{1/2} \geq \theta$ and $\lambda_i \geq \theta$, so

$$\begin{aligned} \mathbb{P} \left(|X * \tilde{X}(k)| \geq 2\theta \left(\sum_{n \in E} (\mathbf{Var} X_n)^2 \right)^{1/2} \right) &\leq \sum_{i=1}^2 \mathbb{P} \left(\left| \sum_{n \in E_i} Y_n \right| \geq \lambda_i \sigma_i \right) \\ &\leq 4 \max(e^{-\theta^2/4}, e^{-\theta/2}). \end{aligned}$$

Since this holds for each $k \neq 0$ and $|\text{supp } X * \tilde{X}| \leq |E|^2$, the conclusion follows (replacing 2θ with θ). \square

Corollary 4.2. *Let $\nu_j^{(\omega)}$ be the random measure defined as before, $0 < \alpha < 1/2$ and $\kappa > 0$. Then there is a set $\Omega_3 \subset \Omega_2$ with $\mathbb{P}(\Omega_3 = 1)$ such that for each $\omega \in \Omega_3$,*

$$(4.2) \quad |\nu_j^{(\omega)} * \tilde{\nu}_j^{(\omega)}(x)| \leq C_\omega \beta(2^j)^{-1} \delta_0(x) + C_\omega \beta(2^j)^{-2} 2^{\kappa j} \sqrt{\sum_{n=1}^{2^j} \tau_n^2}.$$

Proof. For the bound at 0, we use the fact that

$$\begin{aligned} \nu_j^{(\omega)} * \tilde{\nu}_j^{(\omega)}(0) &= \beta(2^j)^{-2} \sum_{n=1}^{2^j} (\xi_n(\omega) - \tau_n)^2 \\ &= \beta(2^j)^{-2} \sum_{n=1}^{2^j} (\tau_n^2(1 - \xi_n(\omega)) + (1 - \tau_n)^2 \xi_n(\omega)) \\ &\leq \beta(2^j)^{-2} \sum_{n=1}^{2^j} (\tau_n + \xi_n(\omega)) \\ &= 2\beta(2^j)^{-1} + \beta(2^j)^{-2} \sum_{n=1}^{2^j} (\xi_n(\omega) - \tau_n) \end{aligned}$$

so that

$$\mathbb{P}(\nu_j^{(\omega)} * \tilde{\nu}_j^{(\omega)}(0) > 3\beta(2^j)^{-1}) \leq \mathbb{P}\left(\sum_{n=1}^{2^j} (\xi_n(\omega) - \tau_n) > \beta(2^j)\right) \leq 2 \exp(-\frac{1}{2}\beta(2^j))$$

for j sufficiently large, by Chernoff's inequality. The Borel-Cantelli Lemma implies that $\nu_j^{(\omega)} * \tilde{\nu}_j^{(\omega)}(0) \leq 3\beta(2^j)^{-1}$ for j sufficiently large (depending on ω), so there exists C_ω with $0 \leq \nu_j^{(\omega)} * \tilde{\nu}_j^{(\omega)}(0) \leq C_\omega \beta(2^j)^{-1}$ for all j .

For the other term, we note that $\mathbf{Var} \xi_n \leq \tau_n$, so we set $\theta = 2^{\kappa j}$ and apply Lemma 4.1:

$$\mathbb{P}\left(\beta(2^j)^2 \|\nu_j^{(\omega)} * \tilde{\nu}_j^{(\omega)}\|_{\ell^\infty(\mathbb{Z}^\times)} \geq 2^{\kappa j} \left(\sum_{n=1}^{2^j} \tau_n^2\right)^{1/2}\right) \leq 4 \cdot 2^{2j} \exp(-2^{\kappa j}/4)$$

which sum over j . The Borel-Cantelli Lemma again proves the bound holds with probability 1. \square

Note that $\sum_{n=1}^{2^j} \tau_n^2 = \Theta(2^{(1-2\alpha)j})$; thus for $\alpha < 1/2$ and $\kappa + \epsilon = 1/2 - \alpha$,

$$\beta(2^j)^{-2} \left(\sum_{n=1}^{2^j} \tau_n^2\right)^{1/2} 2^{\kappa j} = O(2^{(-\frac{3}{2} + \alpha + \kappa)j}) = O(R_j^{-(1+\epsilon)}).$$

Therefore the measures $\nu_j^{(\omega)}$ satisfy the bound (3.1), for all $\omega \in \Omega_3$. Since $\mu_j^{(\omega)} - \nu_j^{(\omega)} = \mathbb{E}\mu_j$ is a weighted average of the regular ergodic averages, $\sup_j |\varphi * \mathbb{E}\mu_j| \leq C \sup_N |\varphi * N^{-1} \chi[1, N]|$ so that Birkhoff's Ergodic Theorem implies the needed weak ℓ^1 bound; and the ℓ^∞ maximal inequality for $\mu_j^{(\omega)}$ is trivial. Thus Proposition 3.1 implies the weak maximal inequality (2.2), and we have proved Theorem 1.1.

Remark This argument does not require τ_n to obey a power law. If τ_n is decreasing

and if $\beta(2^j)^{-2} \sqrt{\sum_{n=1}^{2^j} \tau_n^2} \leq C2^{-(1+\epsilon)j}$ for some $\epsilon > 0, C < \infty$ and all j , the sequence $\{n : \xi_n(\omega) = 1\}$ will be universally L^1 -good with probability 1.

It remains, finally, to note that $\{n : \xi_n = 1\}$ indeed has Banach density 0 (with probability 1) if the τ_n decrease more rapidly than some power law. Conveniently enough, a converse result also holds:

Proposition 4.3. *Let $\{\tau_n\}$ be a decreasing sequence of probabilities, and let ξ_n be independent Bernoulli random variables with $\mathbb{P}(\xi_k = 1) = k^{-\alpha}$. Then if $\tau_n = O(n^{-\alpha})$ for some $\alpha > 0$, the sequence of integers $\{n : \xi_n = 1\}$ has Banach density 0 with probability 1 in Ω ; otherwise, it has Banach density 1 with probability 1 in Ω .*

Proof. It is elementary to show that

$$(4.3) \quad 2^{-r} \tau_{r(n+1)}^m \leq \mathbb{P} \left(\sum_{j=rn}^{r(n+1)-1} \xi_j \geq m \right) \leq 2^r \tau_{rn}^m.$$

(We majorize or minorize the ξ_j by i.i.d. Bernoulli variables and use the Binomial Theorem.) Then if $\tau_n = O(n^{-\alpha})$, let $K > 0$ and fix $m, r \in \mathbb{N}$ such that $m\alpha > 1$ and $r > mK$; the probabilities above are then summable, so the first Borel-Cantelli Lemma implies that on a set Ω_K of probability 1 in Ω , there exists an M_ω such that for all $n \geq M_\omega$, $\sum_{j=rn}^{r(n+1)-1} \xi_j < m < \frac{r}{K}$; then it is clear that $\{n : \xi_n = 1\}$ has Banach density less than $3K^{-1}$. Let $\Omega' = \bigcap_K \Omega_K$; then $\mathbb{P}(\Omega') = 1$ and $\{n : \xi_n = 1\}$ has Banach density 0 on Ω' .

For the other implication, note that if $\tau_n \neq O(n^{-1/R})$, there exists a sequence n_k with $n_{k+1} \geq 2n_k$ such that $\tau_{n_k} \geq n_k^{-1/R}$; then

$$\sum_{n=1}^{\infty} \tau_{Rn}^R \geq R^{-1} \sum_{n=2}^{\infty} \tau_n^R \geq R^{-1} \sum_{k=2}^{\infty} (n_k - n_{k-1}) \tau_{n_k}^R \geq R^{-1} \sum_{k=2}^{\infty} \frac{1}{2} = \infty.$$

Thus the probabilities in (4.3) are not summable in n , for $m = r = R$. Since the variables ξ_n are independent, the second Borel-Cantelli Lemma implies that there is a set $\tilde{\Omega}_R$ of probability 1 on which $\{n : \xi_n(\omega) = 1\}$ contains infinitely many blocks of R consecutive integers. Therefore if $\tau(n) \neq O(n^{-\alpha})$ for every $\alpha > 0$, let $\tilde{\Omega}' = \bigcap_R \tilde{\Omega}_R$; on this set of probability 1, $\{n : \xi_n = 1\}$ has Banach density 1. \square

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